Xing Liu February 13, 2023

Imperial College London

Main Reference for the Talk

A. Anastasiou, A. Barp, F.-X. Briol, et al. (2021) Stein's Method Meets Statistics: A Review of Some Recent Developments

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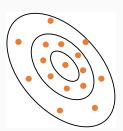
- 1. Motivation
- 2. Kernelized Stein Discrepancy
- 3. Application 1: Goodness-of-Fit Testing
- 4. Application 2: Sample Quality Quantification
- 5. Application 3: Sample Approximation

Motivation

Let Q, P be probability measures on $\mathcal{X} \subset \mathbb{R}^d$.

- P admits a density $p = p^*/Z$, where Z is an unknown normalising constant.
- \bullet Samples are observed from Q only.

Problem of interest: How to quantify the discrepancy between P and another probability measure Q?



P: target distribution
Q: MCMC samples

0000	0000
1//1	1111
2222	2222
3333	3333
9444	4486
5555	5555
6626	6668
7777	デファク
8823	8888
9999	9999

P: a generative model Q: true images

Integral Probability Metrics (IPM)¹

Given a family $\mathcal{H} \subset L^1(P) \cap L^1(Q)$ of real-valued functions, the IPM is:

$$d_{\mathcal{H}}(Q, P) = \sup_{h \in \mathcal{H}} |\mathbb{E}_{X \sim Q}[h(X)] - \mathbb{E}_{X \sim P}[h(X)]|.$$

- Total Variation distance: $\mathcal{H} = \{h : \mathcal{X} \to \mathbb{R} : \sup_{x \in \mathcal{X}} |h(x)| \le 1\}$
- L^1 -Wasserstein distance: d_W : $\mathcal{H}_W = \{h : \mathcal{X} \to \mathbb{R} : |h(x) - h(y)| \le ||x - y||_2, \forall x, y\}$
- Bounded Wasserstein distance/Dudley metric: d_{bW} : $\mathcal{H}_{bw} = \{h \in \mathcal{H}_W : h \text{ is bounded}\}$

Problem: $d_{\mathcal{H}}(Q, P)$ requires integrating over P, so it cannot be computed

Solution: Choose \mathcal{H} so that $\forall h \in \mathcal{H}$, $\mathbb{E}_{X \sim P}[h(X)] = 0$.

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How to choose \mathcal{H} for a generic P? — Use Stein's method!

4

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Stein's Method

Given a probability measure P on \mathcal{X} , we are interested in finding a linear operator \mathcal{T} acting on some set $\mathcal{G}(\mathcal{T})$ of functions on \mathcal{X} such that

Stein's Identity

For any probability measure Q on \mathcal{X} ,

$$Q = P \iff \mathbb{E}_{X \sim Q}[(\mathcal{T}g)(X)] = 0, \text{ for all } g \in \mathcal{G}(\mathcal{T}).$$
 (1)

Glossary:

- Stein operator: \mathcal{T}
- Stein class: $\mathcal{G}(\mathcal{T})$ for which $\mathbb{E}_{X \sim Q}[(\mathcal{T}g)(X)] = 0$ for all $g \in \mathcal{G}(\mathcal{T}g)$
- Stein set: Any $\mathcal{G} \subset \mathcal{G}(\mathcal{T})$
- Stein characterisation: The equivalence (1)



Charles Stein

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Setup: P,Q two probability measures. P has unnormalised density p that is continuously differentiable.

Recall: The IPM is
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Kernelized Stein Discrepancy Given a Stein operator
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 and a Stein set $\mathcal G$, the Stein discrepancy is:
$$\mathbb S(Q,P,\mathcal G)=\sup_{g\in\{\mathcal Tg\colon g\in\mathcal G\}}\|\mathbb E_{X\sim Q}[(\mathcal Tg)(X)]\|_2.$$

Ideally, we want

- Separation: $S(Q, P, \mathcal{G}) = 0 \iff Q = P$
- Computability: $\mathbb{S}(Q, P, \mathcal{G})$ can be efficiently computed even when the normalising constant of p is unknown and sampling from P is infeasible.

How to choose \mathcal{T} ? Langevin Stein operator

$$(\mathcal{T}g)(x) = \langle \nabla \log p(x), g(x) \rangle + \langle \nabla, g(x) \rangle.$$

How to choose \mathcal{G} ? Reproducing Kernel Hilbert Spaces (RKHS)

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How to choose \mathcal{G} ? Reproducing Kernel Hilbert Spaces (RKHS)!

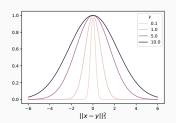
Reproducing Kernel Hilbert Spaces (RKHS)

Reproducing kernel: $k: \mathcal{X} \times \mathcal{X} \to \mathbb{R}$.

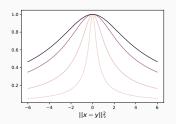
- Symmetric: k(x, y) = k(y, x).
- Positive definite: For any $n \in \mathbb{Z}_+$, $x_1, \ldots, x_n \in \mathcal{X}$ and $c_1, \ldots, c_n \in \mathbb{R}$, $\sum_{i,j=1}^n c_i c_j k(x_i, x_j) \ge 0$.

RKHS: A Hilbert space \mathcal{H}_k is a RKHS associated with k if

- $\forall x \in \mathcal{X}, k(\cdot, x) \in \mathcal{H}.$
- Reproducing property: $\forall x \in \mathcal{X}, \forall f \in \mathcal{H}, \langle f, k(\cdot, x) \rangle_{\mathcal{H}_k} = f(x).$



Radial basis function (RBF): $k(x,y) = \exp\left(-\frac{1}{\gamma}||x-y||_2^2\right)$



Inverse multi-quadric (IMQ): $k(x,y) = \left(1 + \frac{1}{\gamma} ||x - y||_2^2\right)^{-1/2}$

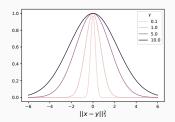
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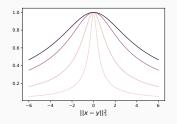
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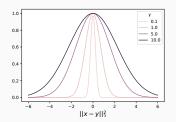
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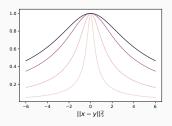
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(Langevin) Kernelized Stein Discrepancy (KSD)²

Choosing $\mathcal{G}_k^d := \times_{j=1}^d \mathcal{G}_k$ for $\mathcal{G}_k := \text{unit-ball in a RKHS } \mathcal{H}_k$, the KSD is

$$\mathbb{D}(Q, P) := \mathbb{S}^2(Q, P, \mathcal{G}_k^d) = \mathbb{E}_{X, X' \sim Q}[k_P(X, X')],$$

where

$$\begin{aligned} \mathbf{k}_{P}(x, x') &\coloneqq k(x, x') \langle \mathbf{s}_{p}(x), \mathbf{s}_{p}(x') \rangle + \langle \nabla_{x} k(x, x'), \mathbf{s}_{p}(x') \rangle \\ &+ \langle \nabla_{x'} k(x, x'), \mathbf{s}_{p}(x) \rangle + \langle \nabla_{x}, \nabla_{x'} k(x, x') \rangle, \end{aligned}$$

and
$$s_p(x) := \nabla_x \log p(x)$$
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- k_P : Stein reproducing kernel.
- $\mathbb{D}(Q, P) \ge 0$ and $\mathbb{D}(Q, P) = 0 \iff Q = P$.
- k_P is computable even if p is only known up to a normalisation: $s_p(x) = \nabla_x \log p(x) = \nabla_x \log(p^*(x)/Z) = \nabla_x \log p^*(x) - \nabla_x Z$.
- Estimation: given i.i.d. $\{X_i\}_{i=1}^n \sim Q$,

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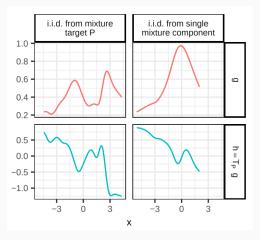


Figure credit: [Gorham and Mackey, 2017]

Setup: P same as before, and $\{Q_n\}_{n\geq 1}$ is a sequence of empirical measure.

Questions

- 1. Does $Q_n \to_d P$ imply $\mathbb{D}(Q_n, P) \to \mathbb{D}(P, P) = 0$?
- 2. Does $\mathbb{D}(Q_n, P) \to 0$ imply $Q_n \to_d P$?

Theorem [Gorham and Mackey, 2017]

- 1. If $\nabla \log p$ is Lipschitz and k is twice continuously differentiable, then $d_W(Q_n, P) \to 0 \implies \mathbb{D}(Q_n, P) \to 0$.
- 2. Assume $\nabla \log p$ is distantly dissipative (a relaxation of log-concavity). If either an IMQ kernel is used or $(Q_n)_{n\geq 1}$ is uniformly tight (a tail condition). Then $\mathbb{D}(Q_n,P)\to 0 \implies Q_n\to_d P$.

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Goodness-of-Fit Testing

Given sample $\{X_i\}_{i=1}^n$ drawn independently from Q, test

$$H_0: Q = P \text{ vs. } H_1: Q \neq P.$$
 $\iff H_0: \mathbb{D}(Q, P) = 0 \text{ vs. } H_1: \mathbb{D}(Q, P) \neq 0$

KSD test³: Compute test statistic \mathbb{D}_n using $\{X_i\}_{i=1}^n$, and reject for large values.

Given significance level $\alpha \in (0,1)$, the rejection threshold $\hat{q}_{1-\alpha}$ should satisfy

Type-I error
$$:= \mathbb{P}_P(\hat{\mathbb{D}}_n \ge \hat{q}_{1-\alpha}) \le \alpha$$
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To compute $\hat{q}_{1-\alpha}$, we need to know the distribution of \mathbb{D}_n under H_0

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Goodness-of-Fit Testing

Algorithm (KSD Test)

Given $\{x_i\}_{i=1}^n \sim Q$ and a test level $\alpha > 0$,

1. For $b = 1, \ldots, B$, compute

$$\widehat{\mathrm{KSD}^2}_{k,b} \coloneqq \frac{1}{n^2} \sum_{1 \leq i \neq j \leq n} \epsilon_i^b \epsilon_j^b k_P(x_i, x_j),$$

where $\epsilon_1^b, \dots, \epsilon_n^b$ are i.i.d. Rademacher r.v. in $\{-1, 1\}$.

2. Reject if $\hat{\mathbb{D}}^2 \ge \hat{\gamma}_{\alpha} := (1 - \alpha)$ -quantile of $\{\widehat{\text{KSD}}_{k,b}^2\}_{b=1}^B$.

Example — 1D Gaussian Mixture

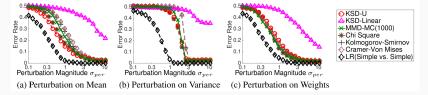


Figure credit: [Liu et al., 2016]

- $P = \sum_{k=1}^{5} w_k \mathcal{N}(\mu_k, \sigma^2)$, where $w_k = \frac{1}{5}$, $\sigma^2 = 1$, and $\mu_k \in [0, 10]$.
- $Q = \text{same as } P \text{ but with Gaussian noise injected into } \mu_k, \sigma^2 \text{ and } \log w_k.$

Blindness of KSD

 $\mathbb{D}(Q,P) \approx 0$ when Q and P are multi-modal distributions with well-separated modes. \longrightarrow KSD test power $\approx \alpha$.

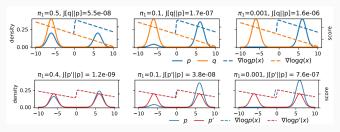


Figure credit: [Wenliang and Kanagawa, 2020]

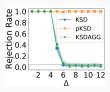


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- Classical diagnostics such as Effective Sample Size and the Gelman-Rubin statistic do not account for asymptotic bias
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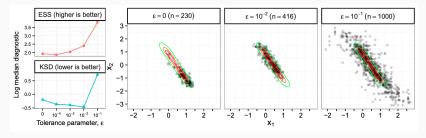
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Example — Hyperparamter Selection



Using KSD to select hyperparameters of a MCMC sampler, with comparisons against ESS (Effective Sample Size). Figure credit: [Gorham and Mackey, 2017]

Application 3: Sample Approximation

Objective: Sampling from P with continuously differentiable density p.

Idea:

- Initialise $X \sim Q$
- Iteratively apply a map $T(X) = X + \epsilon g(X)$ so that $T^{\infty}(X) \sim P$.

Choose g in some function class $\mathcal H$ to maximally decrease $\mathrm{KL}(T_\#Q\|P)$

$$\sup_{g \in \mathcal{H}} \left\{ -\frac{d}{d\epsilon} \mathrm{KL}(T_{\#}Q \| P)|_{\epsilon=0} \right\} = \sup_{g \in \mathcal{H}} \mathbb{E}[T_{g}(T_{g})]$$
 (*)

[Liu and Wang, 2016]:

- (*) = $S(Q, P, \mathcal{H})$, the Stein discrepancy objective!
- Hence, the optimal g^* is the maximiser in (*)
- Choosing $\mathcal H$ to be a RKHS, g^* has an analytical form:

$$g^*(\cdot) = \mathbb{E}_{X \sim Q}[\mathcal{T}k(\cdot, x)] = \mathbb{E}_{X \sim Q}[\underbrace{\nabla \log p(X)k(X, \cdot)}_{attraction} + \underbrace{\nabla_x k(X, \cdot)}_{repulsion}]$$

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Stein Variational Gradient Descent

- Given $X_1, \ldots, X_n \sim Q$ i.i.d., and $\epsilon > 0$.
- For t = 1, 2, ..., set

$$X_i^{(t)} = X_i^{(t-1)} + \frac{\epsilon}{n} \sum_{j=1}^n k(X_i^{(t)}, X_j^{(t)}) \nabla \log p(X_j^{(t)}) + \nabla_X k(X_i^{(t)}, X_j^{(t)}) .$$

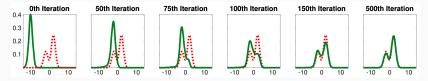


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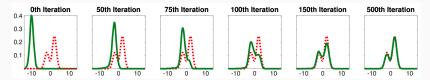


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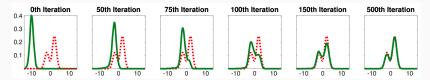


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Variance Collapse

Even in moderate dimensions, SVGD particles will collapse onto the modes of P and exhibit no diversity.

 https://github.com/ImperialCollegeLondon/GSVGD/blob/main/ imgs/gsvgd_cover.gif

Solutions:

• Work on low-dim projected spaces: [Gong et al., 2021a, Gong et al., 2021b, Liu et al., 2022b].

Yet There are Many More...

- Post-processing of MCMC samples [Riabiz et al., 2020].
- Stein points [Chen et al., 2018, Chen et al., 2019].
- Model training [Barp et al., 2019, Grathwohl et al., 2020].
- ..

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